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DEVELOPMENT OF FINANCIAL MARKETS AND ECONOMIC GROWTH: CASE OF POLAND AND IRELAND

Abstract. Impact of financial markets development on economic growth is one of the controversial problems in the theory of economics. Results of the R. Levine and others economist's research seems to confirm statistically important connections between financial markets development and economic growth. Polish financial markets have relatively short history (since 1990).

The aim of the paper will be to answer the question if the impact of financial markets development on economic growth in Poland and Ireland existed in the year of 1994–2007 (in Ireland – 1996–2007) and how was strong. In the paper was also presented selected methods and results of other authors research. The author has presented results of the own research based on the econometric analyses of available data concern economies of mentioned above countries.

Key words: economic growth, financial development, financial markets, financial structure, stocks, bonds, econometric model, ordinary least square.

1. INTRODUCTION

The research which has resulted in this paper aimed at answering the following questions:

- 1) What are the relationships between financial development and economic growth from the point of view of the theory of economics and finance?
- 2) Does the development of the financial market in Poland stimulate economic growth in comparison with Ireland and, if yes, in what way?
- 3) How strong is the influence of the financial market development on economic growth in Poland in comparison with Ireland?

The research included an analysis of relationships between financial market development and economic growth in Poland and Ireland in the years 1994–2007 on the basis of the financial development indicators possible to use. This research took advantage of a simple multi-equation model taking into account conclusions drawn from the methodology of examining this type of relationship applied by R.G. King and R. Levine.

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2. FINANCIAL MARKET DEVELOPMENT AND ECONOMIC GROWTH – A REVIEW OF THEORETICAL ASPECTS

Financial market performs crucial functions in economy, such as: ensuring liquidity in economy, appropriation of financial claims, mobilisation of part of national income above current expenditure for investment purposes, strengthening of the motivating role of profit, allocation of capital in economy, economic shock absorption through the risk sharing mechanism.

It is also noteworthy that the role of financial markets in economy depends on the financial structure of economy. The financial structure of a given country's economy is shaped by: institutions, financial technology, rules of the game, which determine how financial activities are organised at a point in time. R. Stultz uses an analogy that „ (...) financial structure is to financial system what a foundation is to a house. Many different houses can be built on the same foundation. However, at the same time, a foundation makes it impossible to build some types of houses. If the foundation is designed for a one floor house, it cannot be used to build a skyscraper” (Stultz R. [2004], pp. 146–147). Thus, the financial systems functioning in different countries can be divided into two categories: market-oriented financial systems, referred to as Anglo-Saxon ones, and bank-oriented financial systems, also known as the continental ones. The former ones and their segments are strongly competitive in relationship to the banking sector as an alternative to capital allocation and raising capital. Financial market plays the main role in capital allocation. Analytical companies connected with the financial market provide information to the entire market. The financial market facilitates complex risk management as the signals received from the market allow investors to assess risk and profitability of investments and enterprises. It also facilitates takeovers and mergers of enterprises which, on the one hand, lead to capital concentration and, on the other hand, exert pressure on managers to work effectively and achieve high profitability of enterprises and investment projects (Allen F., Gale D. [2001]). In the bank-oriented system, banks play the main role. They collect information about enterprises and managers and on the basis of an analysis they allocate capital, enable management of different types of risk and in this way affect effectiveness of investment projects in economy. E.R. Siri and P. Tufano point out the role of banks in capital mobilisation in order to finance ventures leading to attainment of the benefit of scale (Siri E.R., Tufano P. [1985], pp. 81–128).

Several important factors affect the shape of the financial system model. The level of economic development is the first one. Countries of growing GDP per capita tend to evolve towards a more distinct role of the financial market. An important role is played by a tendency to risk on the side of economic entities (enterprises and households) of the system which is culturally conditioned. The

choice of particular forms of financing by enterprises significantly affects the development of the financial market and/or the banking system. One cannot ignore effectiveness of the legal system which regulates financial market functioning and protects shareholders (see: Osiński J., Wyczański P., Tymoczko D., Grąt A. [2004], pp. 15) .

It should be emphasized that at present in many countries financial systems combine, to a different degree, the elements of both models. Ireland represents some of the countries where financial systems are definitely market-oriented and where the ratio of stock-exchange capitalisation to GDP is higher than the ratio of bank credit to GDP. Poland represents bank –oriented financial system.

J. Schumpeter, pointed out specific functions of financial intermediation and financial markets, essential for economic growth and development and consisting in mobilisation of savings, capital allocation, risk management, facilitating transactions, and company monitoring (Schumpeter J. [1960], pp. 72–85, 155–202). Taking into account Schumpeter’s theory of entrepreneur and innovation one can propose a thesis that also in the case of financial institutions and financial intermediation a process of „creative destruction” occurs which results in financial development being a component of economic development. R. Levine referred to J. Schumpeter’s concept. According to him “Financial development occurs when financial instruments, markets, and intermediaries ameliorate – though do not necessarily eliminate – the effects of information, enforcement, and transactions costs and therefore do a correspondingly better job at providing the five financial functions:

- production of ex ante information about possible investments and capital allocation,
- monitoring of investments and exert of corporate governance after providing finance,
- facilitating trading of financial instruments, risk diversification, management of risk,
- mobilizing and pooling of savings,
- ease the exchange of goods and services.

Each of these financial functions may influence savings and investment decisions and hence economic growth” (Levine R. [2004], pp. 5–7).

As the above fragment indicates financial development means, first of all, changes of qualitative character. These functions are implemented by financial markets and financial intermediaries. In the long run implementation of these functions leads to an increase in capital accumulation. Furthermore, through creating possibilities of risk diversification and creating finance sources, they stimulate increase in technological innovations. Together, they stimulate economic growth. J. Greenwood and B. Jovanovich indicated parallels and interdependencies in financial market development and economic growth. According to

them, economic growth provides means thanks to which financial markets and financial intermediation develop. In turn, this process accelerates economic growth by supporting capital allocation (see: Greenwood J., Jovanovic B. [1989], p. 25).

Moreover, the long-term relationships between the financial market development and economic growth is also worth mentioning. The financial market development is a component of a broadly understood financial development. Financial development is expressed, on the one hand, by growing values of specific quantity variables characterising changes in the financial market and in the banking system which embrace, among others, an increase in the number of banks per 1,000 inhabitants, growth in bank assets in relation to GDP, an increase in the bank credit value in relation to GDP, an increase in stock exchange capitalisation in relation to GDP, an increase in the number of public partnerships (whose shares are listed at the stock exchange), an increase in the number of new emissions of financial instruments. On the other hand, one can speak about financial development when certain quality changes occur, like, e.g. launching new bank products and financial innovations, a tendency of economic units to invest savings in new financial products and use new financial services, emergence of new specialist financial institutions meeting new needs of the market, combining and permeating of hitherto separate types of financial activities and financial services (e.g. bank assurance), consolidation of financial institutions and their internationalisation.

According to R. Levine the influence of financial development on economic growth follows the block diagram presented in Figure 1.

It is worth mentioning that the correlations between financial development and, in particular, financial markets and economic growth, is not unambiguous from the theoretical point of view as well as from the point of view of empirical studies based on different econometric methods. J. Robinson claimed that „where enterprise leads, finance follows” (see: Robinson J. [1952], p. 80). R. Lucas also questioned the relationship between financial development and economic growth, claiming that if there is any such relationship, the role of finance in economic growth is overestimated (see: R. Lucas [1988], pp. 3–42).

In recent years there were several publications which were highly skeptical about the impact of financial development on economic growth. These include publications by, among others, P. Wachtel, M.J. Manning and P. L. Rousseau Wachtel P. [2003]; (see: Manning M.J. [2003]; Rosseau, P.L., Wachtel P. [2005]).

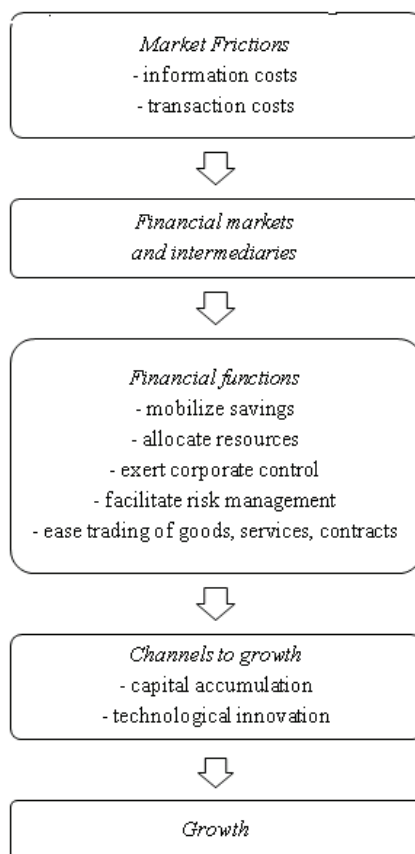


Figure 1. Theoretical relationship between finance and economic growth
Source: (Levine R. [1997], p. 691).

3. REVIEW OF METHODS AND RESULTS OF SELECTED EMPIRICAL INVESTIGATIONS

Majority of empirical research into relationships between financial development and economic growth is based on panel data and thus concerns not particular countries but their groups.

A basic econometric model used in the research into relationships between financial development and economic growth is R. Levine and R.G. King's version of the R. Barro's model of economic growth regression. This model assumes the following form:

$$Y_{it} = \alpha_0 + \alpha F_{it} + \beta X_{it} + u_{it},$$

where Y_{it} is the growth rate of real GDP per capita in i -th country over the period t , F_{it} – is a financial development index in the i -th country over the period t (the ratio of the financial sector's liquid liabilities to GDP, the ratio of credit for the non-financial private sector to GDP, the ratio of credit for the non-financial private sector to total domestic credits, the ratio of domestic assets of deposit banks to domestic assets of the entire banking sector), X_{it} – the vector of basic, predetermined instrumental variables explaining economic growth in the i -th country over the period of t (the natural logarithm of initial GDP per capita, the natural logarithm of the scholarisation index – the ratio of children registered in secondary schools to the total number of children at the school age, the share of foreign trade turnover in GDP, the ratio of government consumption to GDP, the ratio of budget deficit to GDP). The model is estimated with the use of the double ordinary least square (2OLS) method.

After GDP per capita has been replaced in the equation with the per capita capital growth rate, and next with the per capita efficiency growth rate, and the investment rate in GDP – the authors investigated the influence of financial development indicators on these quantities using the same model form.

It is also worth mentioning that R.G. King and R. Levine used the statistical panel data based on the World Bank's statistics – World Development Indicators and partly those of IFS – International Financial Statistics of the International Monetary Fund. On the other hand, P.L. Rousseau and P. Wachtel used the World Bank's database – World Development Indicators. R. King and R. Levine carried out research into relationships between the figures of real GDP per capita and the size of financial intermediation measured by the ratio of the financial system's liquid liabilities to GDP, based on a sample consisting of 80 countries and covering the 1960–1989 period. Next, they investigated the influence of financial development indicators on long-term economic growth rates per capita, capital accumulation and productivity growth. In each case the correlation indicators were high and statistically significant but different depending on a group of countries – divided into countries of low, medium and high economic development (see: King R.G., Levine R. [1993], pp. 717–737). Further research based on panel data also confirmed a relatively strong impact of financial development, including that of financial markets, on economic growth (see: Levine R., Loyaza N., Beck T. [2000], ss. 31–77; Caporale G.M., Howells P.G.A., Soliman A.M. [2005], pp. 166–176).

On the other hand, P.L. Rousseau and P. Wachtel's research did not fully confirm the results obtained by the above mentioned authors. P.L. Rousseau and P. Wachtel applied the same research method as R. Levine and R.G. King, and used in their research the annual panel data comprising statistical data from 84 countries for the 1960–2003 period (Rousseau P.L., Wachtel P. [2005]).

4. RESULTS OF EMPIRICAL RESEARCH INTO RELATIONSHIP BETWEEN FINANCIAL MARKET DEVELOPMENT AND ECONOMIC GROWTH IN POLAND AND IRELAND IN THE YEARS 1994–2007

4.1. Statistical data and indicators

The research used statistical data comprised in the database „*Financial structure dataset (Nov. 2008)*” developed by T. Beck and E. Al.-Hussainy according to the methodology described in „*A New Database on Financial Development and Structure*” by T. Beck, A. Demiriuguç – Kunt and R. Levine. The statistical data used in the research and concerning shaping of financial development indicators are derived entirely from this database. On the other hand, the data referring to GDP per capita and physical capital and investments are derived from the databases of Eurostat and the Central Statistical Office (GUS), as well as the World Development Indicator base of the World Bank . (see: Beck T., Al.-Hussainy E. [2007]; Beck T., Al.-Hussainy E. [2008]; Beck T., Demiriuguç-Kunt A., Levine R. [1999]).

The data referring to Poland cover the years 1994–2007 (14 years) and for Ireland cover the years 1996–2007(12 years). In the case of Ireland, available date for years 1994–1995 are not existing. The author has chosen the above period for his analysis due to the fact that in Poland only since 1994 it has been possible to speak about significant financial development indicators in relation to GDP. Thus, the work covers 14 observations concerning the following quantities: the growth rate of real GDP per capita (*GDPp*), the growth rate of real gross physical capital (*CAPITAL*), efficiency growth rate (*EFF*), the growth rate of real gross investment per capita (*GINV*) and the two financial market development indicators presented below.

A relatively small sample entailed several problems. Due to the restricted number of degrees of freedom in an econometric model, the number of explanatory variables had to be restricted, which, obviously affected the quality of obtained models. Moreover, data inconsistency may have occurred due to the fact that not all needed quantities were available in statistics of one database. Particularly, the data referring to the growth rates of real gross investment per capita and physical capital per capita were taken from other sources than the World Development Indicators of the World Bank, and more specifically from the databases of EUROSTAT and GUS. For example, there is no data concerning physical capital and its growth for Poland and other new EU countries in the World Bank and Eurostat bases. Therefore, the author used here a variable known as “gross fixed assets”, which is available in Statistical Yearbooks published by GUS.

In order to determine the efficiency growth indicator (*EFF*), combining in itself effects of technology use, human resource and labour productivity, the method applied by R.G. King and R. Levine was used (King R.G., Levine R. [1993], s. 722). The starting point is an economic growth equation in the form:

$$y = k^a x,$$

where y stands for real GDP per capita, k represents real physical capital per capita, x stands for remaining determinants of per capita GDP growth (joint factor of technology, human resources and labour), a – is a parameter of production function. This equation can be transformed by logarithming the sides of the equation to the form:

$$\ln y = a \ln k + \ln x.$$

By changing symbols, this dependence can be converted (as R.G. King and R. Levine did it) to the form: $GDPp = aCAPITAL + EFF$, where $GDPp$ represents the growth rate of real GDP per capita, $CAPITAL$ – stands for the growth rate of physical capital per capita, and EFF indicates the efficiency growth rate, a – is a share of physical capital growth in real per capita GDP growth. Hence, EFF can be determined in the following way:

$$EFF = GDPp - aCAPITAL.$$

Following R.G. King and R. Levine, the value $a = 0.3$. Additionally, the growth rate of real gross investment per capita ($GINV$) was used in the study.

Financial market development indicators were constructed in the following way (see: Beck T., Al.-Hussainy E. [2007]; Beck T., Al.-Hussainy E. [2008]; Beck T., Demircic-Kunt A., Levine R. [1999]):

$STOCK$ – the ratio of the stock-exchange capitalization on the stock market to GDP $\{0,5[STOCK_t/P_{e_t} + STOCK_{t-1}/P_{e_{t-1}}]\}/GDP_t/P_{a_t}$, $PBONDS$ – the ratio of the stock-exchange capitalization on the bonds market to GDP $\{0,5[PBONDS_t/P_{e_t} + PBONDS_{t-1}/P_{e_{t-1}}]\}/GDP_t/P_{a_t}$.

In the above formulas the following symbols were adopted: P_{e_t} – stands for an inflation rate (CPI) at the end of the year, $P_{e_{t-1}}$ is an inflation rate (CPI) at the beginning of the year, P_{a_t} – represents the average yearly inflation rate (CPI), GDP_t – GDP in the year t , $t - year$.

4.2. Research method

In the study an econometric multi-equation model (simple model) was used. The model consists of 3 equations, each of which is estimated separately by the OLS method. These equations were constructed in such a way as to show the effect of both the real sphere and the financial sphere on economic growth, or on its major factors: capital growth and efficiency growth. The choice of the model was dictated by the following premises:

- Basic exogenous and endogenous models of growth indicate the following economic growth factors: physical capital, human capital, labour, technology; for this reason the dependencies between per capita real GDP growth and per capita physical capital growth (*CAPITAL*) and efficiency growth (*EFF*) constituting a joint effect of labour, human capital and technology productivity growth were adopted.
- Conclusions drawn from the financial development theory point out that individual components of financial development affect economic growth by creation of possibilities to accumulate capital (physical capital growth) and technological innovation (*EFF* growth), which lead to economic growth.
- A tendency to examine relationships between the real and financial spheres of economy, and, in particular, between the indicators characterizing financial market development and economic growth and its factors – physical capital, efficiency and investment.
- A relatively short time sequence limiting the number of explanatory variables due to the required number of degrees of freedom as well as limiting possibilities of stylized fact identification.

The first equation characterizes the correlation between per capita real GDP growth and its main factors:

$$(1) \ln_GDPp = \ln_a_{10} + a_{11} \ln_EFF + a_{12} \ln_CAPITAL + u_1,$$

where \ln_GDPp – represents the logarithm of per capita GDP growth, \ln_EFF – stands for the logarithm of efficiency growth rate, $\ln_CAPITAL$ – is the logarithm of the growth rate of capital per capita, \ln_a_{10} – stands for the logarithm of the constant.

The second equation characterizes the correlation between the growth rate of real GDP per capita and efficiency growth (*EFF*) and the ratios of the stock market capitalization to GDP and the bond-market capitalization to GDP (*PBONDS*).

$$(2) GDPp = a_{20} + a_{21}EFF + a_{22}STOCK + a_{23}PBONDS + u_2$$

The third equation characterizes the relationship between the growth rate of real physical capital per capita (*CAPITAL*) and the growth rate of gross investment, and the ratios of the stock market capitalization to GDP (*STOCK*) and bond market capitalization to GDP (*PBONDS*).

$$(3) \text{ CAPITAL} = a_{30} + a_{31}GINV + a_{32}STOCK + a_{33}PBONDS + u_3$$

4.3. Presentation of research results

The first equation reveals the obvious significant stochastic relationships between GDP per capita and explanatory (independent) variables being the economic growth factors. The relationship between per capita GDP growth and per capita physical capital growth and efficiency growth is obvious from the point of view of economic growth models (see: Tables 1A,B). Due to the non-linear dependence (power function), the author used logarithms of the per capita real GDP growth rate and logarithms of the growth rates of real physical capital per capita and efficiency.

Table 1A. Poland. Equation 1: OLS Method estimation with the use of 14 observations 1994–2007
Dependant variable: l_GDPp

	Coefficient	Standard error	t-Student	p value	
Const	-0.154094	0.227312	-0.6779	0.51184	
l_CAPITAL	0.204385	0.0646612	3.1609	0.00906	***
l_EFF	0.658898	0.0214507	30.7169	<0.00001	***
Mean of dependent variable	-3.161880	Standard deviation of dependent variable		0.549790	
Sum of squared residuals	0.041929	Standard error of residuals		0.061739	
Unadjusted R-squared	0.989330	Adjusted R- squared		0.987390	
F(2, 11)	509.9452	p value for the F-test		1.43e-11	
Log – Likelihood	20.81067	Durbin-Watson statistics		2.541708	
Autocorrelation of Residuals – rho1	-0.303210				

*** the variable is significant at the significance level of 0.01,

** the variable is significant at the significance level of 0.05,

* the variable is significant at the significance level of 0.1.

Source: the author's own calculations with the use of the GRETL program.

Table 1B. Ireland. Equation 1: OLS Method estimation with the use of 14 observations 1994–2007
Dependent variable: l_GDPp

	<i>Coefficient</i>	<i>Standard terror</i>	<i>t-Student</i>	<i>p value</i>	
Const	0.142907	0.0692364	2.0640	0.06342	*
l_EFF	0.843363	0.0105352	80.0517	<0.00001	***
l_CAPITAL	0.121753	0.0163836	7.4314	0.00001	***
Mean of dependent variable		-2.695873	Standard deviation of dependent variable		0.344828
Sum of squared residuals		0.002611	Standard error of residuals		0.015405
Unadjusted R-squared		0.998311	Adjusted R- squared		0.998004
F(2, 11)		3251.224	p value for the F-test		5.65e-16
Log – Likelihood		40.24567	Durbin-Watson statistics		1.952762
Autocorrelation of Residuals – rho1		-0.050234			

*** the variable is significant at the significance level of 0.01,

** the variable is significant at the significance level of 0.05,

* the variable is significant at the significance level of 0.1

Source: the author's own calculations with the use of the GRETL program.

In the case of Poland Equation 2 indicates a statistically significant positive relationship between real GDP growth per capita and the ratio of stock-market capitalization to GDP. An increase in stock-market capitalization by 1 percentage point causes an increase in per capita GDP growth rate by 0.021 percentage point. On the other hand, the relationship between the ratio of bond-market capitalization to GDP and the growth rate of GDP per capita is negative, hence an increase in the share of the bond-market capitalization in GDP by 1 percentage point causes a decrease in the growth rate of real GDP per capita by 0.051 percentage point (see: Table 2A). The influence of both variables on the growth rate of real GDP per capita is definitely lower than that of efficiency.

Irish economy reveals a statistically significant positive relationship between the ratio of stock-market capitalization to GDP and the growth rate of real GDP per capita. However, the effect of this explanatory variable on the explained variable is weaker than in Poland. An increase in the ratio of stock-market capitalization to GDP by 1 percentage point causes an increase in the growth rate of GDP per capita by 0.0093 percentage point. The relationship between the ratio of bond-market capitalization to GDP and the growth rate of real GDP per capita is negative, but statistically insignificant (see: Table 2B).

Table 2A. Poland. Equation 2: OLS Method estimation with the use of 14 observations 1994–2007
Dependent variable: *GDPp*

	<i>Coefficient</i>	<i>Standard value</i>	<i>t-Student</i>	<i>p value</i>	
<i>Const</i>	0.0159835	0.00161021	9.9263	<0.00001	***
<i>EFF</i>	1.04013	0.0206821	50.2914	<0.00001	***
<i>STOCK</i>	0.0219688	0.0057401	3.8273	0.00333	***
<i>PBONDS</i>	-0.0511506	0.00926045	-5.5236	0.00025	***
Mean of dependent variable	0.047357		Standard deviation of dependent variable	0.019069	
Sum of squared residuals	0.000019		Standard error of residuals	0.001361	
Unadjusted R-squared	0.996079		Adjusted R- squared	0.994902	
F(2, 11)	846.7375		p value for the F-test	2.51e-12	
Log – Likelihood	74.87862		Durbin-Watson statistics	3.209570	
Autocorrelation of Residuals – rho1	-0.629110				

*** the variable is significant at the significance level of 0.01,

** the variable is significant at the significance level of 0.05,

* the variable is significant at the significance level of 0.1

Source: the author's own calculations with the use of the GRETL programme.

Table 2B. Ireland. Equation 2: OLS Method estimation with the use of 12 observations 1996–2007
Dependent variable: *GDPp*

	<i>Coefficient</i>	<i>Standard error</i>	<i>t-Student</i>	<i>p value</i>	
Const	0.00567257	0.00386259	1.4686	0.18013	
EFF	0.998107	0.031766	31.4206	<0.00001	***
STOCK	0.00931736	0.00469812	1.9832	0.08264	*
PBONDS	-0.00144696	0.0124469	-0.1163	0.91032	
Mean of dependent variable	0.070250		Standard deviation of dependent variable	0.025641	
Sum of squared residuals	0.000013		Standard error of residuals	0.001284	
Unadjusted R-squared	0.998177		Adjusted R- squared	0.997493	
F(2, 11)	1459.978		p value for the F-test	2.72e-11	
Log – Likelihood	65.30044		Durbin-Watson statistics	1.409016	
Autocorrelation of Residuals – rho1	0.265218				

*** the variable is significant at the significance level of 0.01,

** the variable is significant at the significance level of 0.05,

* the variable is significant at the significance level of 0.1

Source: the author's own calculations with the use of the GRETL programme.

Stock-market capitalization affects also the growth of real physical capital per capita. In the case of Poland this effect is higher than the influence of the growth rate of real gross investment (see: Table 3A). An increase in the ratio of stock-market capitalization to GDP by 1 percentage point increases the growth rate of real GDP per capita by 0.066 percentage point (0.027 percentage point in the case of the growth rate of real gross investment). A negative impact of the bond-market capitalization to GDP ratio on the growth rate of real physical capital per capita is also significant (an increase by 1 percentage point causes a decrease in the growth rate of GDP per capita by 0.153 percentage point).

Table 3A. Poland. Equation 3: OLS Method estimation with the use of 14 observations 1994–2007
Dependent variable: *CAPITAL*

	<i>Coefficient</i>	<i>Standard error</i>	<i>t-Student</i>	<i>p value</i>	
<i>const</i>	0.053085	0.00519872	10.2112	<0.00001	***
<i>GINV</i>	0.0275522	0.0128545	2.1434	0.05770	*
<i>STOCK</i>	0.0665638	0.0187168	3.5564	0.00521	***
<i>PBONDS</i>	-0.153156	0.0301973	-5.0718	0.00048	***
Mean of dependent variable	0.030564	Standard deviation of dependent variable		0.008354	
Sum of squared residuals	0.000194	Standard error of residuals		0.004407	
Unadjusted R-squared	0.785905	Adjusted R- squared		0.721676	
F(2, 11)	12.23608	p value for the F-test		0.001104	
Log - Likelihood	58.43247	Durbin-Watson statistics		3.084796	
Autocorrelation of Residuals - rho1	-0.569964				

*** the variable is significant at the significance level of 0.01,

** the variable is significant at the significance level of 0.05,

* the variable is significant at the significance level of 0.1

Source: the author's own calculations with the use of the GRETL programme.

In the economy of Ireland, like in Poland, a positive and statistically significant relationship between stock-market capitalization and the growth rate of real physical capital per capita is revealed. What is more, an increase in the stock-market capitalization to GDP ratio has a stronger impact on the growth of real physical capital per capita in Ireland than in Poland. Like in the case of Poland, the relationship between bond-market capitalization and the growth rate of real GDP per capita is negative and statistically significant (see: Table 3B).

Table 3B. Ireland. Equation 3: OLS Method estimation with the use of 12 observations
1996–2007 Dependent variable: *CAPITAL*

	<i>Coefficient</i>	<i>Standard terror</i>	<i>t-Student</i>	<i>p value</i>	
const	0.0225292	0.006179	3.6461	0.00653	***
GINV	0.0850699	0.0241618	3.5208	0.00784	***
STOCK	0.0369639	0.00768274	4.8113	0.00134	***
PBONDS	-0.0701747	0.0218971	-3.2048	0.01252	**
Mean of dependent variable	0.036066		Standard deviation of dependent variable	0.005147	
Sum of squared residuals	0.000057		Standard error of residuals	0.002681	
Unadjusted R-squared	0.802723		Adjusted R- squared	0.728744	
F(2, 11)	10.85071		p value for the F-test	0.003420	
Log - Likelihood	56.46559		Durbin-Watson statistics	2.180682	
Autocorrelation of Residuals - rho1	-0.090415				

*** the variable is significant at the significance level of 0.01,

** the variable is significant at the significance level of 0.05,

* the variable is significant at the significance level of 0.1

Source: the author's own calculations with the use of the GRETL programme.

In the case of Poland and Ireland the ratio of public debt to GDP is relatively low (below 60%) and the relationship between the bond-market capitalization to GDP ratio and economic growth is negative. This phenomenon requires a thorough investigation on a bigger sample and with the use of panel data comprising two groups of countries: countries of high public debt with reference to GDP and countries of a relatively low ratio of public debt to GDP.

5. CONCLUSIONS

An analysis of relationships between selected indicators of financial market development and economic growth in Poland in the period from 1994 to 2007 points to the following stylized facts:

- In the analyzed period, there was a statistically significant relationship between financial market development and economic growth in Poland and Ireland;
- An increase in the government bond-market capitalization (result of increased budget deficit and public debt) has a negative and relatively strong impact on the rate of real economic growth and the rate of real physical capital growth in Poland and Ireland, hence it may confirm a negative effect “portfolio crowding out” caused by increased budget deficit and public debt on economic growth;

- There is a statistically significant but relatively weak, positive relationship between the stock-market capitalization and growth rates of real GDP per capita and real physical capital per capita; this relationship is stronger in the case of Poland than in the case of other examined countries.

To sum up, the conducted analysis revealed a statistically significant and meaningful impact of financial market development on economic growth in Poland and Ireland. The research done is of preliminary character and the conclusions formulated above require further verification by extensive studies covering a larger number of explanatory variables and a larger number of observations.

REFERENCES

- Allen F., Gale D., *Comparing Financial Systems*, The MIT Press, Cambridge, Massachusetts, London 2001.
- Allen F., Santomero A.M., *The Theory of Financial Intermediation*, "Journal of Banking and Finance" 1997 No 21, s. 1461–1486.
- Barro R.J., Sala-i-Martin X., *Economic Growth*, Second Edition, The MIT Press, Cambridge 2004.
- Beck T., Al.-Hussainy E., *Financial Structure dataset, Revised November 2007*, http://siteresources.worldbank.org/INTRES/Resources/469232-07449512766/FinStructure_2006.xls
- Beck T., Al.-Hussainy E., *Financial Structure dataset, Revised November 2008*, http://siteresources.worldbank.org/INTRES/Resources/469232-07449512766/FinStructure_2007.xls
- Beck T., Demirgüç-Kunt A., Levine R., *A New Database on Financial Development and Structure*, Policy Research Working Paper N0 2146, The World Bank Development research Group Finance, July 1999.
- Caporale G. M., Howells P.G.A., Soliman A.M., *Endogenous Growth Models and Stock Market Development: Evidence from Four Countries*, "Review of Development Economics" 2005, Vol. 9, No 2, pp. 166–176.
- Demirgüç-Kunt A., Levine R.(ed.), *Financial Structure and Economic Growth. A Cross-Country Comparison of Banks, Markets, and Development*, The MIT Press, Cambridge Massachusetts, London 2004.
- Fiedor B., *Teoria innowacji. Krytyczna analiza współczesnych koncepcji niemarksistowskich*, Państwowe Wydawnictwo Naukowe, Warszawa 1979.
- Greenwood J., Jovanovic B., *Financial Development, Growth and the Distribution of Income*, NBER "Working Papers, 1989 No 3189.
- King R.G., Levine R., *Finance and Growth. Schumpeter Might Be Right*, "The Quarterly Journal of Economics", 1993, Vol. 108, No 3, August, pp. 7171–737.
- Levine R., *Stock Markets, Growth, and Tax Policy*, "Journal of Finance" 1991, No 46, s. 1445–1465.
- Levine R., *More on Finance and Growth: More Finance, More Growth?*, The Federal Reserve Bank of St. Louis "Review" 2003, Vol. 85, No 4, July/August, pp. 31–46.
- Levine R., *Finance and Growth: Theory and Evidence*, 2004 http://www.econ.brown.edu/fac/Ross_levine/Publication.
- Levine R., *Financial Development and Economic Growth: Views and Agenda*, "Journal of Economic Literature", 1997, Vol/ XXXV (June), pp. 688–726.
- Levine R., Loyaza N., Beck T., *Financial Intermediation and Growth: Casuality and Causes*, "Journal of Monetary Economics", August 2000.
- Lucas R., *On Mechanism of Economic Development*, "Journal of Monetary Economics", 1988, No 22, pp. 3–42.

- Manning M.J., *Finance Causes Growth: Can We Be So Sure?*, *Contributions to Macroeconomics* 3, 2003, www.bepress.com/bejm/contributions/vol3/iss1/art12.
- Neimke M., *Financial Development and Growth in Transition Countries*, Ruhr University Bochum "IEE Working papers" Vol. 173, Bochum 2003.
- Osiński J., Wyczański P., Tymoczko D., Grąt A., *Rozwój systemu finansowego w Polsce w latach 2002–2003*, NBP, Warszawa, grudzień 2004.
- Pietrzak B., Polański Z., *System finansowy w Polsce. Lata dziewięćdziesiąte*, Wydawnictwo Naukowe PWN, Warszawa 2000.
- Robinson J., *The Generalization of the General Theory*, in: *The Rate of Interest and Other Essays*, Macmillan, London 1952.
- Rousseau P.L., Wachtel P., *Economic Growth and Financial Depth: Is the Relationship Extinct Already?*, "WIDER Discussion Paper" 2005, Vol. 2005/10.
- Scholtens B., Wensveen van D., *The Theory of Financial Intermediation: An Essay on What It Does (Not) Explain*, SUERF – The European Money and Finance Forum, Vienna 2003.
- Schumpeter J., *Teoria rozwoju gospodarczego*, PWN Warszawa 1960.
- Siri E.R., Tufano P., *The Global Financial System: A Functional Approach*, Harvard Business School Press, Cambridge 1985.
- Stulz R., *Does Financial Structure Matter for Economic Growth? A Corporate Finance Perspectives*, (w:) Demirgüç – Kunt A., Levine R.(ed.), *Financial Structure and Economic Growth. A Cross-Country Comparison of Banks, Markets, and Development*, The MIT Press, Cambridge Massachusetts, London 2004.
- Wachtel P., *How Much Do We Really Know about Growth and Finance?*, Federal Reserve Bank of Atlanta "Economic Review", 2003, First Quarter.

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ROZWÓJ RYNKÓW FINANSOWYCH A WZROST GOSPODARCZY: PRZYPADK POLSKI I IRLANDII

Wpływ rozwoju rynków finansowych na wzrost gospodarczy jest jednym z kontrowersyjnych problemów w teorii ekonomii. Wyniki badań R. Levine'a i innych ekonomistów zajmujących się tym zagadnieniem wydają się potwierdzać statystycznie istotny związek między rozwojem rynków finansowych a wzrostem gospodarczym. Celem referatu jest odpowiedź na pytanie czy występuje statystycznie istotny związek między rozwojem rynków finansowych a wzrostem gospodarczym w przypadku Polski i Irlandii. Badaniami objęto w przypadku Polski okres 1994–2007, zaś w przypadku Irlandii okres 1996–2007. Wybór okresu wynikał z dostępności danych statystycznych, dotyczących badanego zjawiska w obu krajach. W referacie zaprezentowano wyniki wybranych badań innych autorów oraz wyniki badań własnych autora.

Słowa kluczowe: wzrost gospodarczy, rozwój finansowy, rynki finansowe, struktura finansowa, akcje, obligacje, model ekonometryczny, klasyczna metoda najmniejszych kwadratów.